

Clinical evaluation of maxillary sinus floor elevation with or without bone grafts: a systematic review and meta-analysis of randomised controlled trials with trial sequential analysis

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Abstract

Introduction: Our goal was to systematically review the current evidence comparing the relative effectiveness of two maxillary sinus floor elevation (MSFE) approaches (internal and external) without bone grafts with that of conventional/grafted MSFE in patients undergoing implantation in the posterior maxilla.

Material and methods: Medical databases (PubMed/Medline, Embase, Web of Science, and Cochrane Library) were searched for randomised controlled trials published between January 1980 and May 2023. A manual search of implant-related journals was also performed. Studies published in English that reported the clinical outcomes of MSFE with or without bone material were included. The risk of bias was assessed using the Cochrane Handbook Risk Assessment Tool. Meta-analyses and trial sequence analyses were performed on the included trials. Meta-regression analysis was performed using pre-selected covariates to account for substantial heterogeneity. The certainty of evidence for clinical outcomes was assessed using GRADEpro GDT online (Guideline Development Tool).

Results: Seventeen studies, including 547 sinuses and 696 implants, were pooled for the meta-analysis. The meta-analysis showed no statistically significant difference between MSFE without bone grafts and conventional MSFE in terms of the implant survival rate in the short term ($n = 11$, $I^2 = 0\%$, risk difference (RD): 0.03, 95% confidence intervals (CI): $-0.01-0.07$, $p = 0.17$, required information size (RIS) = 307). Although conventional MSFE had a higher endo-sinus bone gain ($n = 13$, $I^2 = 89\%$, weighted mean difference (WMD): -1.24 , 95% CI: $-1.91- -0.57$, $p = 0.0003$, RIS = 461), this was not a determining factor in implant survival. No difference in perforation ($n = 13$, $I^2 = 0\%$, RD = 0.03, 95% CI: $-0.02-0.09$, $p = 0.99$, RIS = 223) and marginal bone loss ($n = 4$, $I^2 = 0\%$, WMD = 0.05, 95% CI: $-0.14-0.23$, $p = 0.62$, no RIS) was detected between the two groups using meta-analysis. The pooled results of the implant stability quotient between the two groups were not robust on sensitivity analysis. Because of the limited studies reporting on the visual analogue scale, surgical time, treatment costs, and bone density, qualitative analysis was conducted for these outcomes.

Conclusions: This systematic review revealed that both non-graft and grafted MSFE had high implant survival rates. Owing to the moderate strength of the evidence and short-term follow-up, the results should be interpreted with caution.

Key words: maxillary sinus floor elevation, dental implants, bone grafts, meta-analysis, trial sequential analysis.

Introduction

According to a previous observational research, the first teeth to be lost due to periodontal disease are usually the maxillary first and second molars [1]. After the loss of maxillary molars, the residual bone height is insufficient owing to the alveolar bone atrophy and maxillary sinus pneumatization [2]. Bone deficiency is regarded as a challenge for operators in maintaining the stable osseointegration of implants locally. In 1986, Tatum first proposed the lateral window osteotomy as a technique for bone augmentation in the posterior maxilla [3]. Boyne and James used autogenous marrow and bone to raise the Schneiderian membranes and promote osteogenesis [4]. The disadvantages of this technique include postoperative swelling, a long healing period, and the need for numerous operations [5]. In order to improve the inadequacy, Summer proposed a new technique called transalveolar maxillary sinus floor elevation, which used an osteotome to make a 'green-stick' fracture prior to placement of implants [6]. Compared to lateral window osteotomy, the advantages of this technique are minimal invasiveness, fewer postoperative reactions, and shorter operation time. However, the indication for surgery is a minimum residual bone height, ranging from 4 to 6 mm, to achieve the primary stability of the implants [7]. With improvements in implant surfaces, acceptable primary stability has been achieved in areas of bone deficiency [8]. In the past, most researchers considered it necessary for maxillary sinus floor elevation (MSFE) to be combined with graft materials. Currently, MSFE using graft materials is unnecessary. Lundgren *et al.* first reported new bone formation in an area without bone grafts after removing a cyst from the maxillary sinus [9]. Riben *et al.* conducted retrospective studies of patients who underwent MSFE without bone grafts. The results showed a high survival rate of 94.3%, with an average endo-sinus bone gain (ESBG) of 6 mm and an acceptable implant stability quotient of 77 [10]. MSFE without bone grafts prevents maxillary sinusitis caused by membrane perforation, which leads to bone graft leakage. However, this means a lower economic burden for patients in terms of MSFE [11].

According to the present systematic review that reported MSFE without bone grafts, the results demonstrated high survival rates for implants placed in the posterior maxilla [12–14]. However, fewer systematic review have simultaneously included randomised controlled trials (RCTs) on transalveolar maxillary sinus floor elevation and lateral window osteotomy for non-graft MSFE. At the same time, our study performed sample size estimation for the clinical outcomes in the meta-analyses.

The goal of this review was to use evidence-based medicine methods to evaluate the clinical outcomes of MSFE without graft materials based on current evidence.

Material and methods

This systematic review was conducted in accordance with the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) 2020 Statement [15]. To minimise selection bias and avoid duplication, the protocol for this review was registered in the International Prospective Register of Systematic Reviews (PROSPERO) on 20 May, 2023 (CRD42023424526; available from www.crd.york.ac.uk).

Search strategy

An electronic search of four databases (PubMed/Medline, Embase, Web of Science, and Cochrane Library) for studies published in English between January 1980 and May 2023 was conducted. MeSH and entry terms were combined to search for potential studies. The search string on EMBASE was as follows: ('sinus floor augmentation' OR 'augmentations, sinus floor' OR 'floor augmentation, sinus' OR 'floor augmentations, sinus' OR 'sinus floor augmentations' OR 'maxillary sinus floor augmentation' OR 'sinus augmentation therapy' OR 'augmentation therapies, sinus' OR 'augmentation therapy, sinus' OR 'sinus augmentation therapies' OR 'therapies, sinus augmentation' OR 'therapy, sinus augmentation') AND ('Bone Substitutes' OR 'Replacement Material, Bone' OR 'Replacement Materials, Bone' OR 'Materials, Bone Replacement' OR 'Bone Substitute' OR 'Substitute, Bone' OR 'Substitutes, Bone' OR 'Bone Replacement Material' OR 'Material, Bone Replacement' OR 'Bone Replacement Materials').

A manual search of implant-related journals was also performed from January 1980 until May 2023, including *Clinical Oral Implants Research*, *Clinical Implant Dentistry and Related Research*, *International Journal of Implantology*, *Implant Dentistry*, *Journal of Periodontal and Implant Science*, *Journal of Oral Implantology*, *Implantologie*, *International Journal of Oral & Maxillofacial Implants*. A further search of the references of implant-related reviews was conducted for literature that could not be retrieved from electronic databases and journals. A single examiner (J.C.) performed the search.

Eligibility criteria

For potential articles to be included in the meta-analyses, the following criteria had to be met: (1) Participants – Patients with a loss of bone height in the posterior maxilla underwent

treatment for maxillary sinus floor elevation during implant surgery. In addition, all patients were in good general health.

- (2) Interventions – Patients without sufficient residual bone height in the posterior maxilla underwent graftless maxillary sinus floor elevation.
- (3) Comparators – Patients without sufficient residual bone height in the posterior maxilla undergoing maxillary sinus floor elevation with bone grafts.
- (4) Outcomes – (1) survival rate of implants, (2) perforation, (3) ESBG, (4) stability of implants, (5) bone density (BD), (6) marginal bone loss (MBL) height, (7) surgical time (ST), and (8) treatment cost.
- (5) Studies – Clinical trials on humans.

Exclusion criteria were as follows: (1) Animal studies, (2) *in vitro* studies, (3) conferences, (4) patents, (5) one-arm studies, (6) studies presenting incomplete data, (7) non-RCTs, (8) low-quality studies, and (9) review articles.

Data extraction

All electronically retrieved and manually searched studies were imported into the EndNote X9 software. Data were extracted from articles initially retrieved by two independent observers (J.C. and Y.L.). The demographic data of the included articles were recorded on a spreadsheet, including the time of publication, region, number of people, age, sinus sample size, implant sites, residual bone height, approaches used, types of implants, kinds of grafting materials, and follow-up durations. Survival rate as the main outcome was extracted by screening full texts. Additional outcomes, if there were any, recorded for included studies were perforation, ESBG, BD, implant stability quotient (ISQ), MBL, ST, visual analogue scale (VAS), and treatment costs. If there was any disagreement concerning inclusion or exclusion in the collection procedure, another author (J.X.) was consulted. If any data were incomplete, the original authors were emailed for further details.

Risk of bias assessment and certainty of evidence

RCTs included in the quantitative analysis were evaluated for risk of bias by two reviewers (J.C. and Y.L.) as part of data collection. Any divergence was resolved by consensus. The criteria used to assess the quality of the included RCTs were adapted from the guidelines proposed in the Cochrane Handbook, which provided items for the following parameters: sequence generation, allocation concealment, blinding of participants and personal information, blinding of outcome data, incom-

plete outcome data, selective reporting, and other biases [16]. The risk of bias was explicitly judged in each criterion as “low”, “high”, or “unclear”. Allocation concealment, blinding of outcome data, and incomplete outcome data were considered to summarise the evidence quality. Each trial was classified as having a low risk of bias when all three parameters were assessed at a low risk of bias, a high risk of bias when at least one parameter was at a high risk of bias, and a moderate risk of bias in the remaining condition. GRADEpro GDT online was used to evaluate the strength of the evidence for the main and additional outcomes [17]. Each GRADE item (study design, risk of bias, inconsistency, indirectness, imprecision, and other considerations) was assessed for all outcomes, according to the GRADE handbook. Simultaneously, the importance of the data outcomes was judged on a 9-point scale based on clinical issues.

Statistical analysis

Statistical analysis was performed using Review Manager 5.4 and Stata MP 17. The risk difference (RD) with 95% confidence intervals (CIs) was used as the effect size for dichotomous outcomes, as the relative risk (RR) and odds ratio (OR) could not be calculated for zero events. For continuous data, the standard mean difference (SMD) or weighted mean difference (WMD) with 95% CIs was used as the effect measure. If the relative outcomes of the included studies were inadequate, descriptive analyses were performed without statistical calculation. Heterogeneity was explored among the included articles (using the I^2 test and χ^2 test). A fixed effects model was used when there was quantitative evidence of heterogeneity (i.e. $I^2 < 50\%$ and p -value of χ^2 test > 0.05). If the heterogeneity was significant (i.e. $I^2 > 50\%$ and p -value of χ^2 test < 0.05), a random-effects model was applied. For outcomes with significant heterogeneity, meta-regression analyses were used to explore the sources of the heterogeneity. Covariates included approaches for vertical bone augmentation, bone substitutes, follow-up period, year of publication, and risk of bias. To examine the stability of the results drawn from the meta-analysis, we used a statistical method of meta-analysis-based influence analysis. Funnel plots and Egger's test were applied to detect potential publication bias [18]. For each outcome, trial sequential analysis (TSA) was used to judge whether the current RCTs had sufficient statistical power to draw a strong conclusion with TSA 0.9.5.10 Beta Java [19]. Sequential monitoring boundaries, known as required information size (RIS), were used to calculate the sample sizes of the included studies. We set the two-sided type-1 error probability at 5% (alpha boundary) and performed cal-

culations with 80% statistical power, assuming an effect type mean and variance based on low-risk studies for continuous outcomes. The RR reduction (RRR) was given a parameter of 10% for dichotomous data. A Z-curve was obtained from the cumulative evidence provided by existing trials. If this Z-curve crosses the 5% alpha boundary, adequate evidence supporting the intervention has been obtained. The current evidence is sufficient to find no statistical difference between the test and control groups if the Z-curve crosses the futility boundary (inner wedge).

Results

Study selection

The search identified 1390 studies after removing duplicates (Figure 1). A total of 645 studies (116 case reports, 150 review articles, 229 animal studies, 39 *in vitro* studies, 4 conferences, 10 patents, and 97 one-arm studies) were excluded after reading their titles and abstracts. We removed 709 articles after full-text screening of the remaining 745 articles because 702 articles were irrelevant to the present subject and seven articles were published in other languages. The remaining 36 trials were then assessed. Finally, 17 studies were included in this meta-analysis [20–36]. The reasons for excluding these 19 studies are listed in Table I. Altogether, the included studies included 483 patients with an atrophic posterior maxilla.

Characteristics of studies

The characteristics of the included studies are shown in Table II. All studies were RCTs published between 2009 and 2023. The sample size of implants was 696, and 547 sinuses were included in this meta-analyses. Internal maxillary sinus elevation was mentioned in eight studies while nine articles used external maxillary sinus elevation as a method for vertical augmentation. The bone

grafts used in MSFE were as follows: xenograft, allograft, autogenous bone, sponge, nanoparticles, and autogenous bone + xenograft. The most commonly used implant in these trials was Straumann. Table III describes the basic information included in the meta-analysis and the follow-up durations ranging from 1 to 120 months. The conclusions drawn from more than half of the included trials were based on short-term observations.

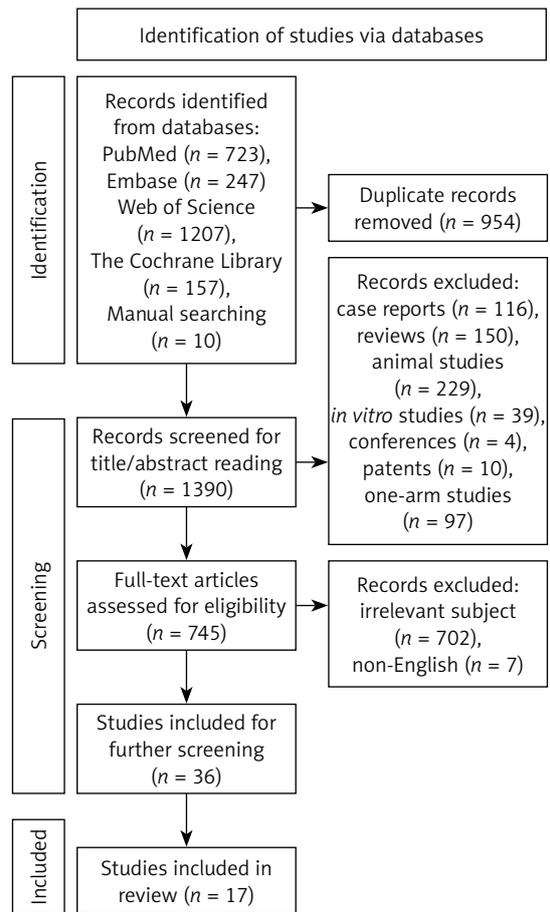


Figure 1. PRISMA 2020 flow diagram for systematic reviews

Table I. Studies excluded from further screening

Author, year	Reasons for exclusion
Aleksa 2016	Data could not be extracted with any instrument
Chitsazi 2018; Nilufer 2015	PRF as an intervention
Diserens 2005; Lundgren 2004; Pommer 2015	One-arm study
Fbris 2020	Retrospective study.
Gabbert 2009; Sina 2010	Without MSFE in control group
Jensen 2023; Loin 2019	Irrelevant outcomes
Johansson 2013; Lie 2019	Histological changes as outcomes
Khalid 2017; Lie 2015; Sohn 2008; Grieson 2013	Inadequate sample size
Mats 2008	Systematic review
Wolfram 2016	MSFE was not mentioned

PRF – platelet-rich fibrin.

Table II. Articles included for review: characteristics

Author, year	Region	Study design	Patients (n)	Sinus (n)	Implants (n)	Implants TG/CG (n/n)	Sinuses TG/CG (n/n)	RBH [mm]	Approach for augmentation	Bone grafts	Type of implant
Felice 2009	Italy	RCTs	10	20	48	24/24	10/10	1-5	External	Xenograft	Geass
Borges 2011	Brazil	RCTs	17	30	54	28/26	15/15	NR	External	AB	NR
Si 2013	China	RCTs	45	45	45	20/21	20/21	2-8	Internal	AB + xenograft	Straumann
Altintas 2013	Turkey	RCTs	14	20	24	12/12	10/10	4-6	External	Allograft	Straumann
Nedir 2013	Switzerland	RCTs	12	19	37	17/20	9/10	0-4	Internal	Xenograft	Straumann
Nedir 2017	Switzerland	RCTs	12	19	37	17/20	9/10	0.9-4	Internal	Xenograft	Straumann
Fouad 2018	Egypt	RCTs	17	20	34	17/17	10/10	4-6	External	Xenograft	NR
Justin 2018	USA	RCTs	33	38	76	36/40	18/20	2-8	External	Allograft	Zimmer
Scarano 2018	Italy	RCTs	23	28	28	14/14	14/14	2-3	External	Xenograft	NR
Khaled 2019	Egypt	RCTs	19	20	25	13/12	10/10	4-6	External	Nanoparticles	Dentis
Trinh 2019	Vietnam	RCTs	30	29	29	15/14	15/14	6-8	Internal	Sponge	Osstem
Qian 2020	China	RCTs	45	45	45	19/21	19/21	2-8	Internal	AB + xenograft	Straumann
Maximiano 2020	Spain	RCTs	49	49	49	25/24	25/24	5-10	Internal	Xenograft	Klockner
Amir 2021	Iran	RCTs	10	18	18	9/9	9/9	NR	External	Xenograft	SIC
Jensen 2022	Denmark	RCTs	39	39	39	19/20	19/20	4-7	External	AB + xenograft	Astra
Zhang 2022	China	RCTs	68	68	68	34/34	34/34	NR	Internal	NR	Osstem
Jensen 2023	Denmark	RCTs	40	40	40	20/20	20/20	NR	Internal	Collagen	Astra

NR – not reported, RBH – residual bone height, AB – autogenous bone, TG – test group (without bone grafts), CG – control group (with bone grafts).

Table III. Articles included for review: outcomes

Author, year	ESBG [mm] Sinus TG/CG	ISQ (value) Implants TG/CG	BD (GV/HU) Sinus TG/CG	MBL [mm] Implants TG/CG	ST [min] Sinus TG/CG	Survival rate (%) Implants TG/CG	Perforation (n) Sinus TG/CG	Follow-up duration
Felice 2009	14.4 (1.7)/ 14.1 (2.5)	NR	NR	NR	19.8 (3.4)/ 20.5 (3.1)	100%/100%	2/1	6 months
Borges 2011	7.91 (3.6)/ 8.31 (2.6)	NR	194.42 (84.75)/ 207 (143.58)	NR	NR	96.4%/100%	1/1	6 months
Si 2013	3.07 (1.68)/ 3.17 (1.95)	NR	NR	1.38 (0.23)/ 1.33 (0.46)	NR	95%/95.2%	2/1	36 months
Altintas 2013	NR	NR	16.25 (438.704)/ 254.91 (246.429)	NR	NR	NR	0/0	6 months
Nedir 2013	3.9 (1.0)/ 5.0 (1.3)	NR	NR	0.6 (0.8)/ 0.4 (0.7)	NR	100%/90%	0/0	12 months
Nedir 2017	3.8 (1.0)/ 4.8 (1.2)	NR	NR	0.6 (0.9)/ 0.7 (1.4)	NR	94.1%/90%	NR	60 months
Fouad 2018	4.85 (0.5)/ 8.59 (0.74)	74 (3.19)/ 78.3 (5.08)	269.08 (16.27)/ 375.598 (49.38)	NR	NR	100%/100%	2/0	6 months
Justin 2018	7.34 (2.47)/ 9.67 (1.64)	81.58 (6.31)/ 83.31 (4.14)	NR	NR	NR	NR	4/5	8 months
Scarano 2018	NR	NR	NR	NR	18.3 (2.1)/ 12.5 (3.1)	NR	2/1	6 months
Khaled 2019	5 (1.5)/ 7 (0.8)	77 (5)/ 78 (5)	420 (23)/ 548 (25)	NR	NR	100%/100%	2/1	6 months
Trinh 2019	1.6 (0.3)/ 3.2 (0.3)	NR	NR	NR	NR	NR	0/0	6 months
Qian 2020	3.14 (1.26)/ 3.07 (1.34)	NR	NR	1.52 (1.08)/ 1.67 (1.06)	NR	95%/90.7%	2/1	120 months
Maximiano 2020	3.28 (1.18)/ 3.47 (1.18)	NR	NR	NR	NR	100%/100%	NR	18 months
Amir 2021	6.20 (2.69)/ 9.56 (4.12)	NR	790.13 (129.03)/ 1094.78 (183.07)	NR	NR	NR	NR	6 months
Jensen 2022	NR	NR	NR	NR	NR	NR	3/4	1 month
Zhang 2022	NR	NR	NR	NR	13.97 (2.42)/ 16.78 (1.72)	96.61%/85%	NR	6 months
Jensen 2023	4.4 (0.7)/ 5.5 (1.2)	82.1 (5.6)/ 80.0 (10.0)	NR	NR	NR	100%/100%	1/0	12 months

Thirteen studies measured ESGB of sinuses as outcomes, four studies measured ISQ of implants, five studies measured BD of sinuses, four studies measured MBL of implants, three studies measured ST of sinuses, 11 studies measured survival rate of implants, and 13 studies measured perforation of sinuses.

Risk of bias assessment

Figures 2 and 3 summarise the risk of bias assessment. Six studies were judged at a low risk of bias and the remaining 11 articles were assessed at a moderate risk of bias. All studies showed a low risk of random sequence generation, allocation concealment, incomplete outcome data, and selective bias. Regarding performance bias, one study reported no blinding of surgeons because the numbered envelopes were opened by the operators on the day of surgery. Ten studies had unclear results regarding this parameter. Nearly half of the studies were unclear regarding the blinding outcome assessment. This may be due to the difficulty in blinding investigators and patients in split-mouth studies. Three studies were judged at a high risk of other biases because these trials were sponsored by an implant company and might have had other potential biases.

Data synthesis and meta-analyses

Survival rate for implants

There were eleven studies that reported data on the survival rate of implants [22–24, 26, 28–32, 34, 36]. Eleven articles demonstrated high survival rates in the without-graft and with-graft groups, with weighted survival rates of 97.93% and 95.21%, respectively. No evidence of heterogeneity was found in this meta-analysis ($p = 0.79, I^2 = 0\%$). A fixed effect was used in the data synthesis. Figure 4 shows no statistical difference between the two groups, with RD of 0.03 (95% CI: -0.01–0.07, $p = 0.17$). The stability of the results is presented in Figure 5. There were no statistical differences after excluding each study individually, suggesting that the results had strong stability. A funnel plot was drawn from studies comparing survival rates (Figure 6). No studies were found outside the funnel plots. This symmetry may indicate the absence of publication bias. Egger's test confirmed this conclusion ($p = 0.062$). Using TSA boundaries to favour the survival rate showed that an RIS of 307 was required to achieve 80% test power. The RIS can be acquired from the pooled data. Meanwhile, the Z-curve crossed the futility boundary and failed to reach the conventional boundaries.

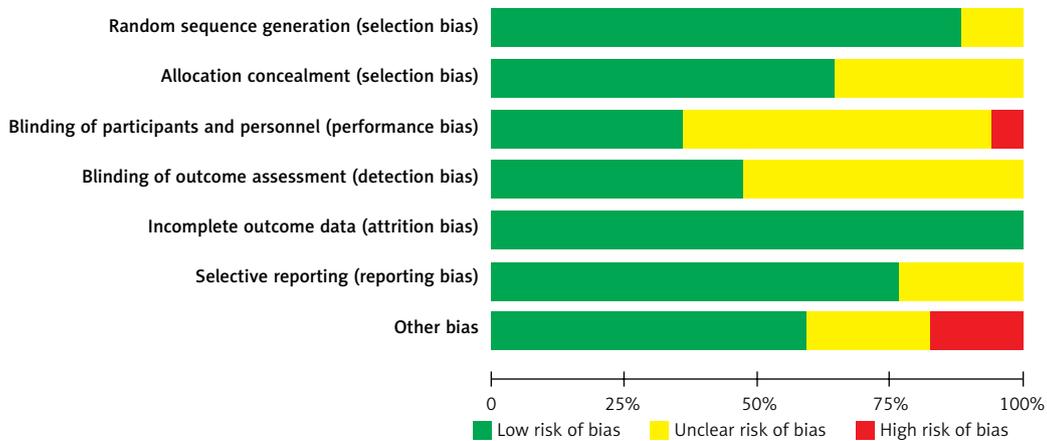


Figure 2. Risk of bias graph

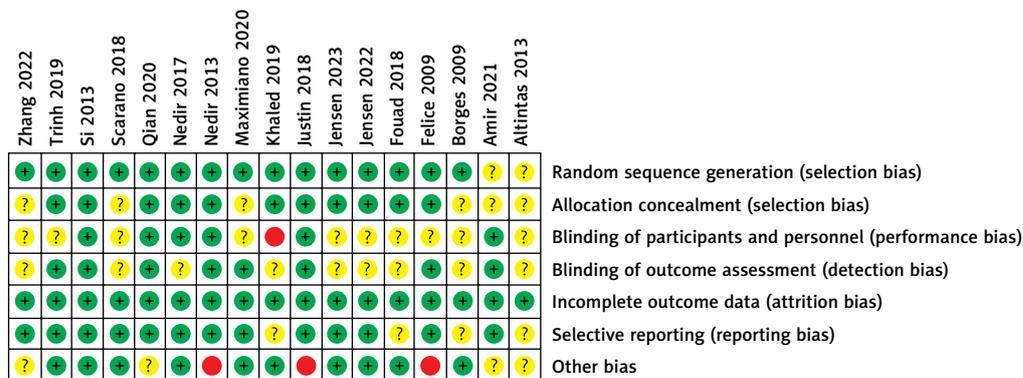


Figure 3. Risk of bias summary

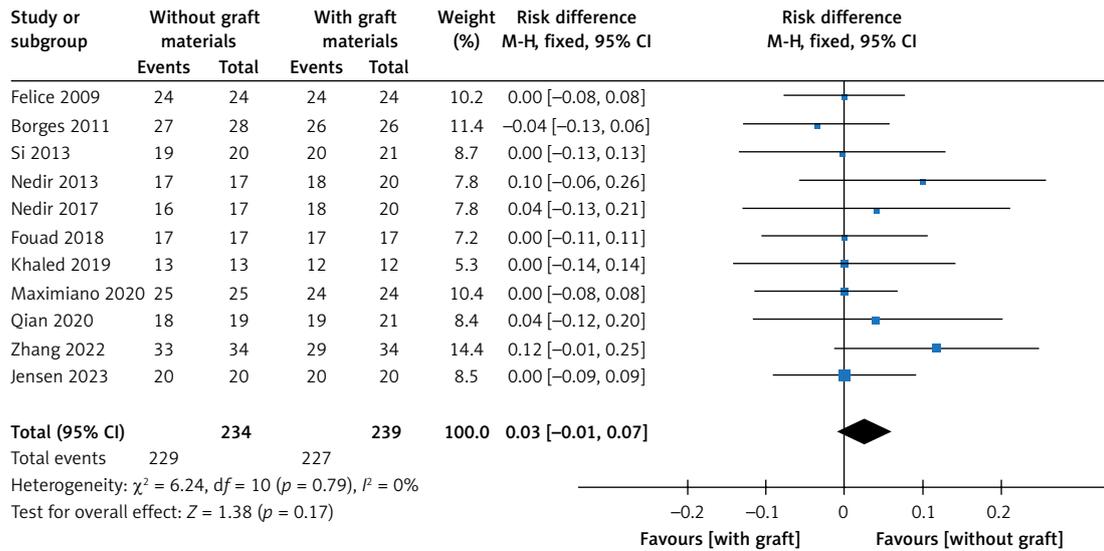


Figure 4. Forest plot comparing survival rates in test group and control group

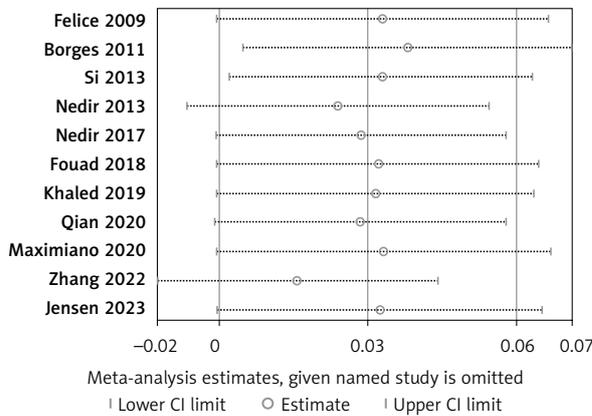


Figure 5. Sensitivity analysis for survival rates

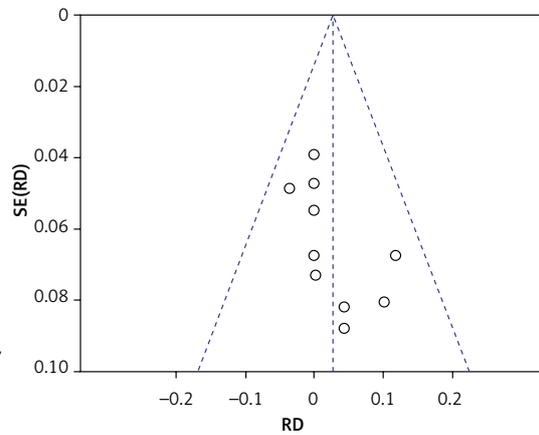


Figure 6. Funnel plot for survival rates

This firmly indicates that the cumulative evidence from the present literature supports no difference between the test and control groups (Figure 7).

Perforation

Thirteen articles described membrane perforation complications during membrane elevation [20, 22–28, 30, 32–35]. A fixed-effects model was applied because there was no evidence of heterogeneity ($p = 0.99$, $I^2 = 0\%$). The pooled outcome showed no statistical difference in perforation between the non-grafted and grafted group (Figure 8): RD = 0.03 (95% CI: -0.02–0.09). The pooled outcomes from the articles were robust in the sensitivity analysis (Figure 9). Funnel plots suggested that no publication bias was present in this meta-analysis (Figure 10). Quantitative analysis of publication bias indicated no small-study effects ($p = 0.306$). The TSA showed firm evidence supporting the lack of difference in perforation between the two methods for MSFE. The Z-curve successfully crossed the futility boundary for per-

foration, and reached an RIS of 223, demonstrating adequate evidence to draw a firm conclusion from the current study (Figure 11).

ESBG

ESBG was measured after at least 6 months in thirteen studies [21–24, 26–32, 34, 35]. A meta-analysis was performed using a random-effects model owing to the high heterogeneity ($I^2 = 89\%$). ESBG was significantly higher in the grafted group, with a WMD of -1.24 (95% CI: -1.91 – -0.57, $p = 0.0003$) (Figure 12). Meta-regression, used to try to explore high between-trial variation, demonstrated that the following covariates had an effect: approaches for vertical bone augmentation (SE = 0.612, $p = 0.044$, covariates: external and internal approach), follow-up period (SE = 0.562, $p = 0.025$, covariates: 6 months and more than 6 months), risk of bias (SE = 0.696, $p = 0.288$, covariates: a low risk of bias and a moderate risk of bias), year of publication (SE = 0.094, $p = 0.309$, covariates: year), and bone substitute (SE = 0.774, $p = 0.085$, covari-

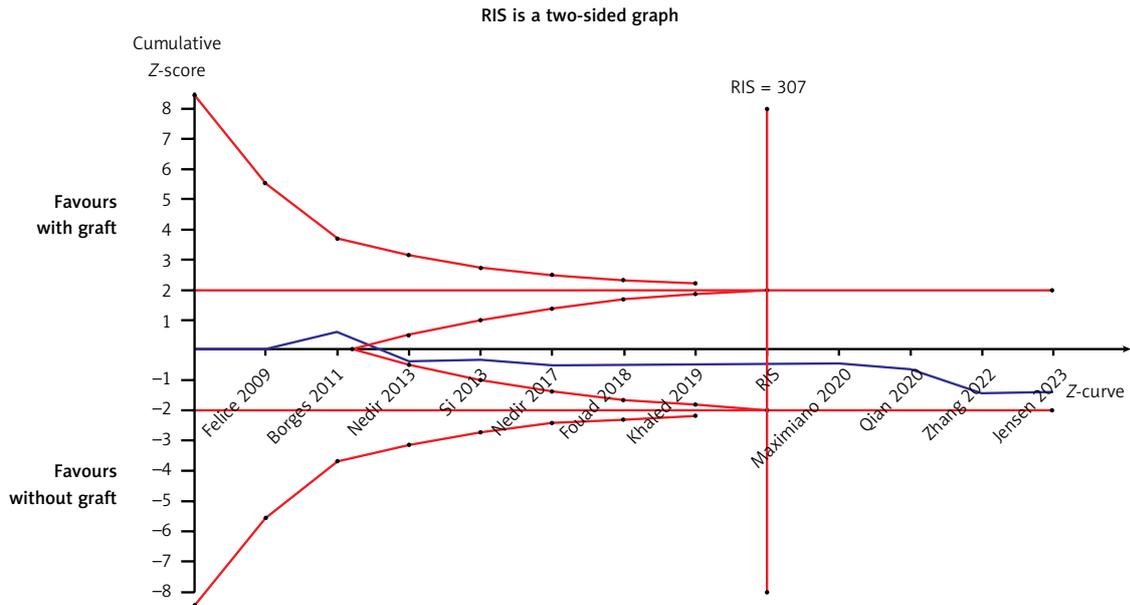


Figure 7. TSA for survival rates

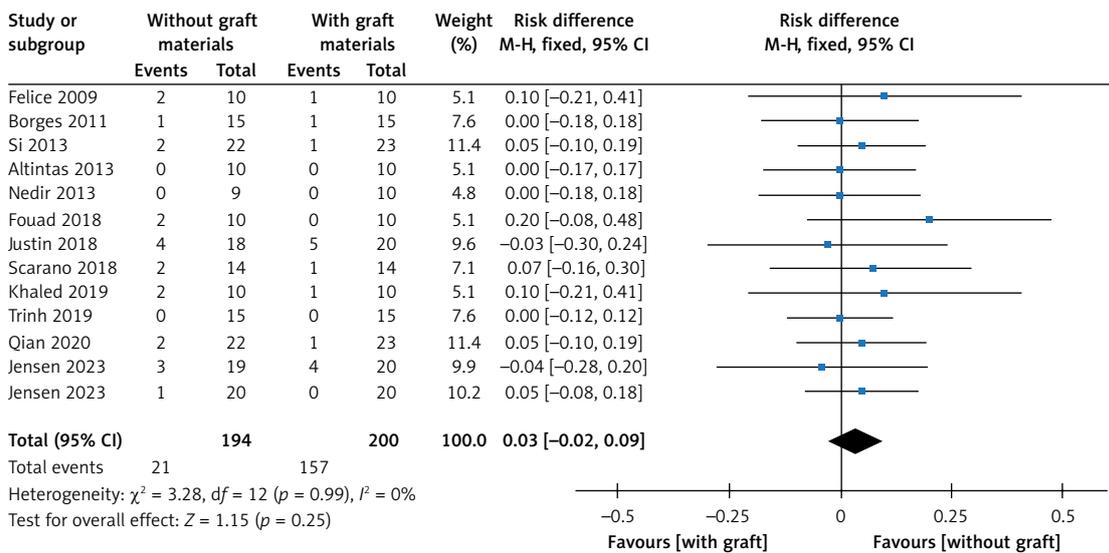


Figure 8. Forest plot comparing perforation in test group and control group

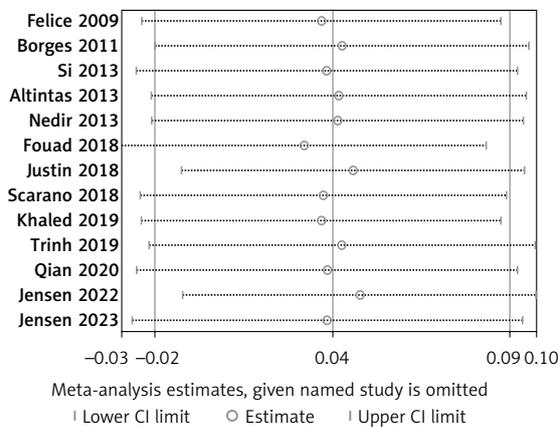


Figure 9. Sensitivity analysis for perforation

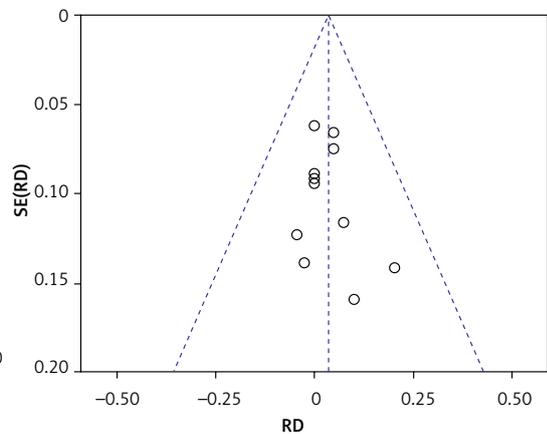


Figure 10. A funnel plot for perforation

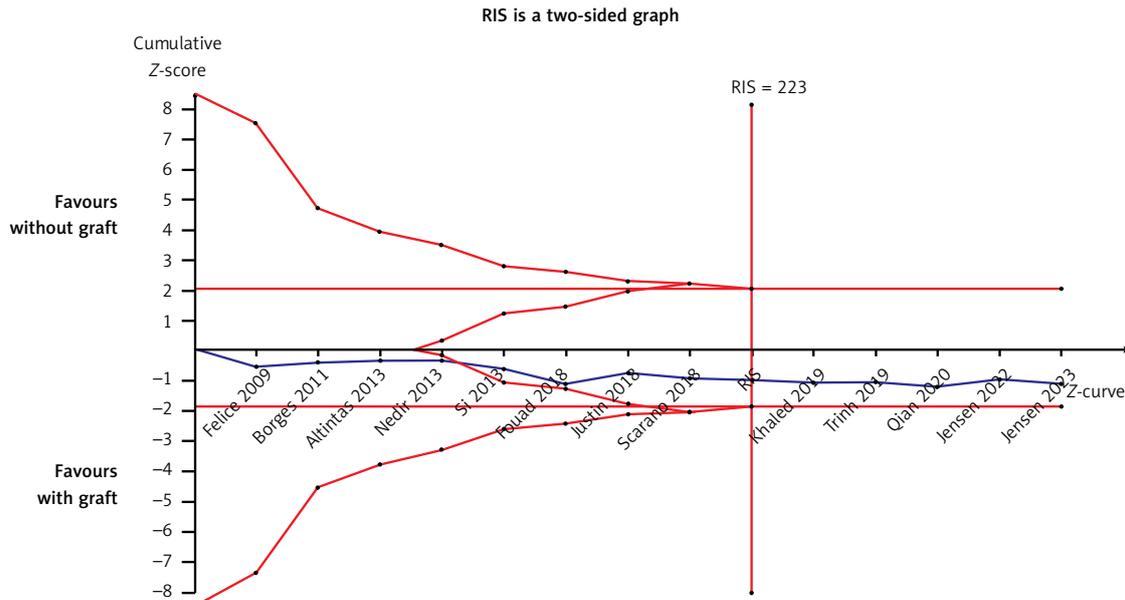


Figure 11. TSA for perforation

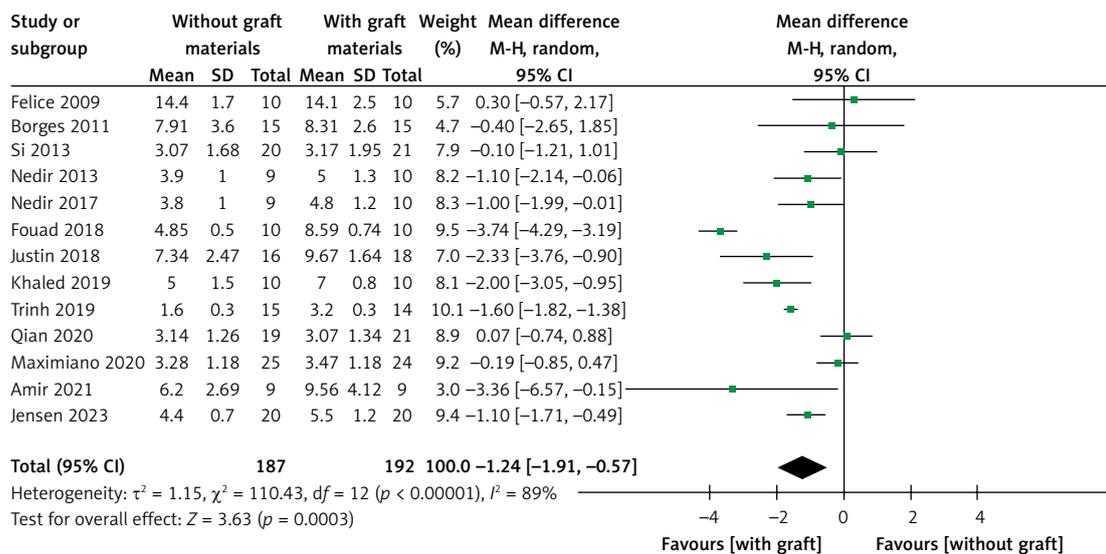


Figure 12. A forest plot comparing ESBG in test group and control group

ates: with AB and without AB). Approaches and follow-up periods may account for between-study heterogeneity. The sensitivity analysis showed that the conclusions had strong stability (Figure 13). The symmetry in Figure 14 and Egger's test ($p = 0.485$) demonstrated no potential publication bias. As shown in Figure 15, the Z curve crossed the conventional and TSA monitoring boundaries although it failed to achieve the expected sample size. The conclusions drawn from the current study are less likely to be false positives.

Implant stability quotient

Four articles compared the ISQ [24, 26–28]. A fixed-effects model was applied in this me-

ta-analysis due to its low heterogeneity ($I^2 = 44\%$). A significant difference was found in ISQ between test and control groups (WMD = -2.04 , 95% CI: -3.66 – -0.41 , $p = 0.01$) (Figure 16). There was no significant difference when Fouad's study [24] was excluded from this meta-analysis (WMD = -0.95 , 95% CI: -2.93 – 1.03 , $p = 0.35$). The Egger's test ($p = 0.281$) and funnel plot indicated no publication bias (Figure 17). The TSA analysis showed that the Z-curve crossed the conventional boundary but failed to cross the TSA monitoring boundary and reached an RIS of 848. The results pooled from the four articles might be due to the presence of false positives, which is consistent with the conclusion from the sensitivity analysis (Figure 18).

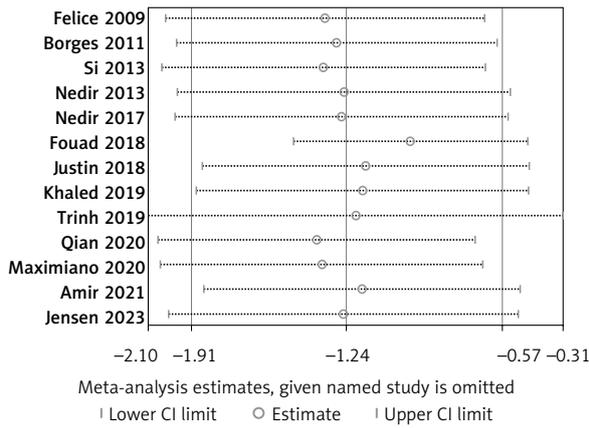


Figure 13. Sensitivity analysis for ESBG

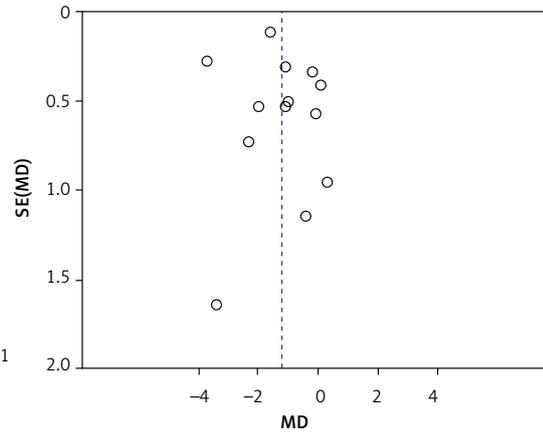


Figure 14. A funnel plot for ESBG

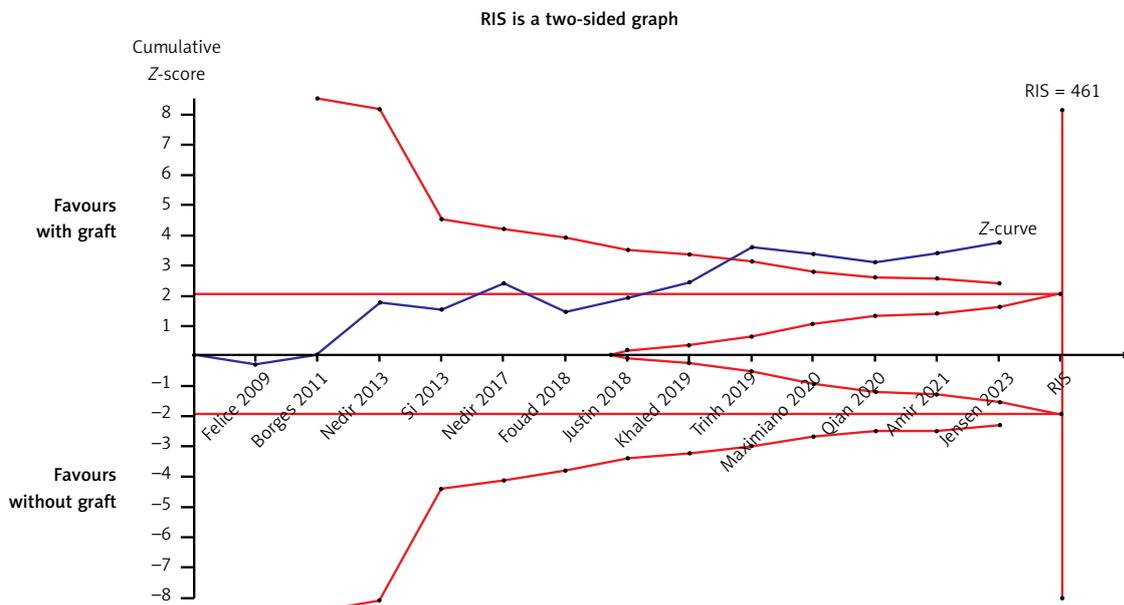


Figure 15. TSA for ESBG

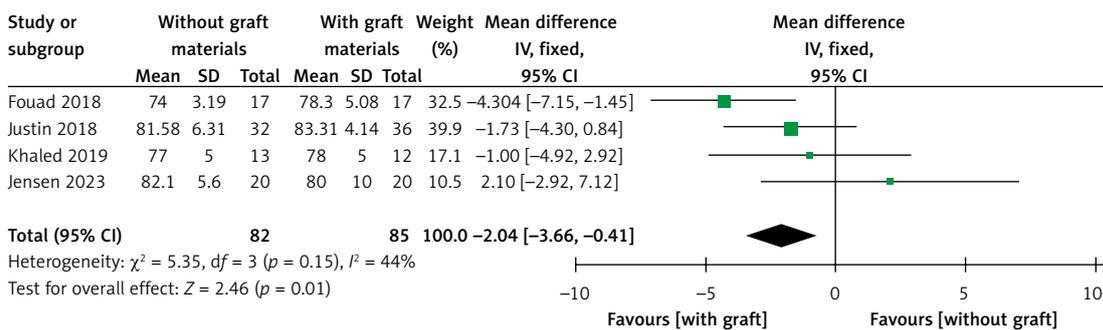


Figure 16. Forest plot comparing ISQ in test group and control group

BD

The data regarding the association between the test and control groups were pooled from five studies [20–22, 24, 28]. The SMD was used in this meta-analysis because of its ability to measure BD in different units. Because of substantial heterogeneity ($I^2 = 88\%$), a random-effects model

was applied for data synthesis. A statistically significant difference was found, with an SMD of -1.90 (95% CI: $-3.29 - -0.51$). A meta-regression analysis was not performed because insufficient studies were included in the meta-analyses. The meta-analysis was abandoned, and a descriptive analysis was used to explore the association be-

tween the two groups. Four articles [20, 21, 24, 28] reported a significant difference between the two groups, with higher BD values in the control group, and only one study [22] indicated that no difference was detected in terms of BD measured 6 months after surgery. AB was used as a bone graft in the control group, compared to the tent group in the Borges study [22]. All articles failed to avoid confounding factors and postmenopausal women with low BD were excluded from the RCTs. The aforementioned parameters may account for partial heterogeneity.

MBL

Four trials with 155 implants reported the MBL results, including three articles with a low risk of bias [30–32, 34]. A fixed effect model was used in this meta-analysis ($I^2 = 0\%$) and no statistically significant difference could be detected from pooled data (WMD = 0.05, 95% CI: -0.14–0.23, $p = 0.62$) (Figure 19). All results from the sensitivity analyses were consistent with those of the prima-

ry analysis (Figure 20). The publication bias was assessed using a funnel plot compiled from the pooled WMD of the included articles (Figure 21). All studies landed in the triangular area of the funnel plot, which was consistent with Egger's test ($p = 0.638$). However, differences were difficult to

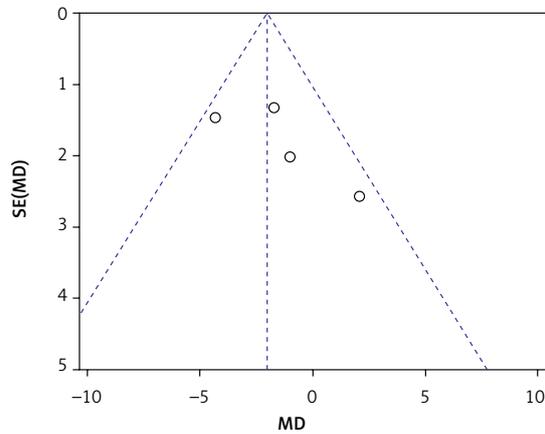


Figure 17. A funnel plot for ISQ

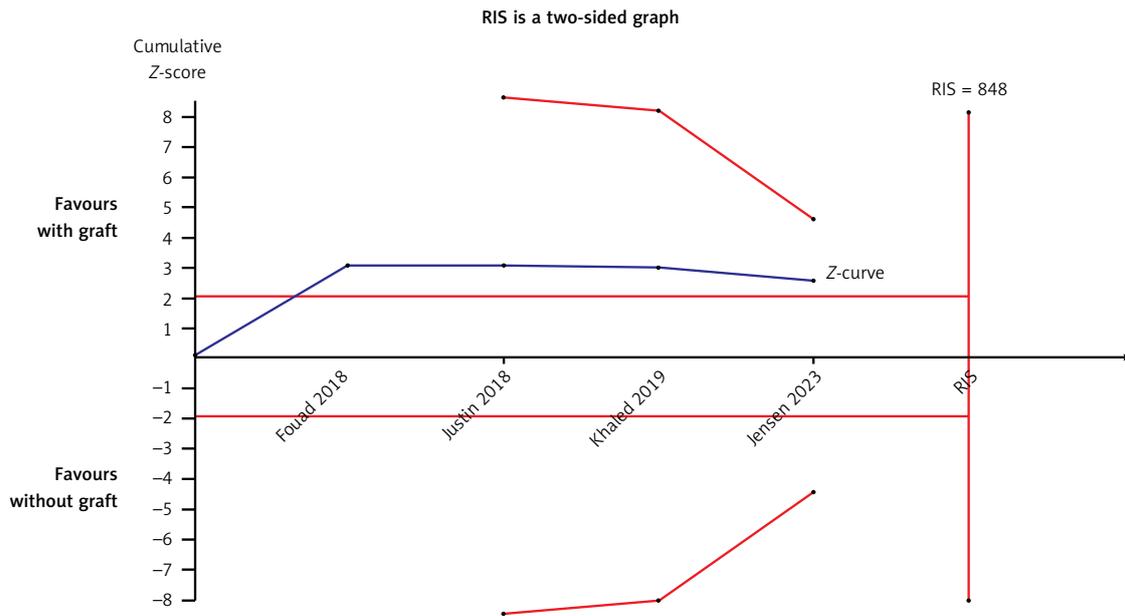


Figure 18. TSA for ISQ

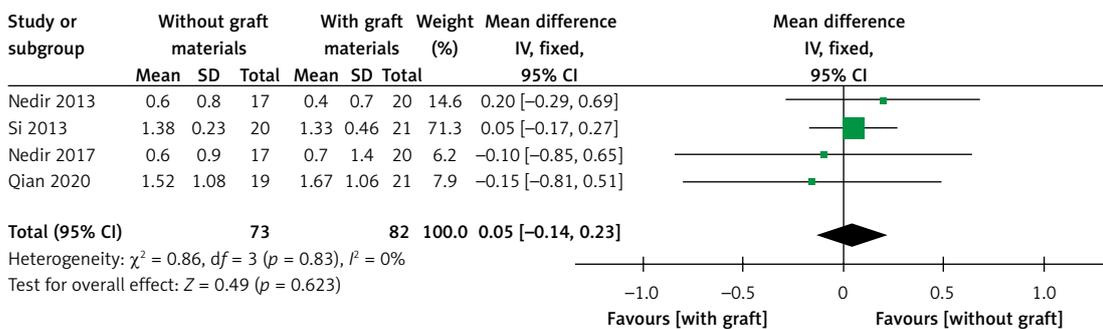


Figure 19. A forest plot comparing MBL in test group and control group

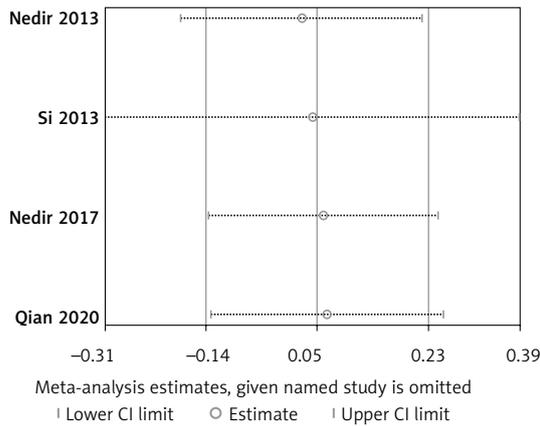


Figure 20. Sensitivity analysis for MBL

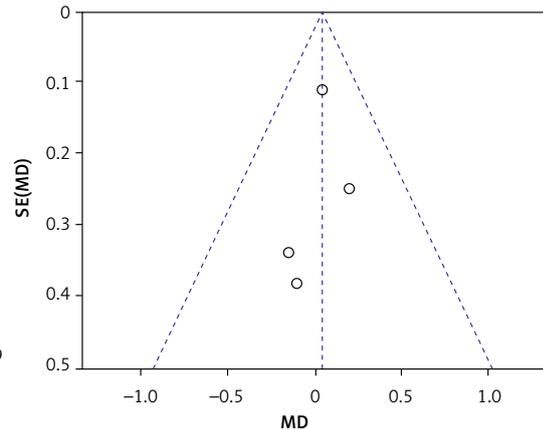


Figure 21. A funnel plot for MBL

detect using Egger’s test because of the inadequate number of included trials. The RIS line could not be computed because of insufficient information.

Surgical time

There were three studies that described the surgical time in the test and control groups [23, 33, 36]. A meta-analysis with a random-effects model ($I^2 = 97$) suggested no difference between the two sides ($p = 0.80$). According to the sensitivity analysis, a statistically significant difference was detected when the study by Scarano *et al.* [33] was excluded from the primary analyses ($p = 0.02$). Due to substantial heterogeneity and insufficient trials, the methods for pooling data were abandoned, and the results of three articles are described individually. Scarano *et al.* found that there was a significant difference in ST required to complete the MSFE procedure: 18.3 (2.1) min for the test group versus 12.5 (3.1) min for the control group. In contrast to Scarano *et al.*, Zhang *et al.* concluded that the mean ST in the test group was 13.97 (2.42) min, which was significantly shorter than that in the control group [36]. However, Felice reported that there was no difference in the ST required to complete the MSFE procedure. In Felice’s study, a rigid barrier was inserted into the sinus, maintaining the sinus lining in the desired position [23].

VAS

A RCT was designed by Jensen in order to explore patient perceptions of recovery after MSFE with or without bone grafts [25]. In the control group, the autogenous bone grafts were harvested from the buccal antrostomy and mixed with the xenograft granules. In the test group, autogenous blood was injected underneath the raised sinus membrane around implants protruding into the maxillary sinus. The research work revealed that there were fewer days of pain and sick leave

in the test group. Therefore, harvesting autogenous bone from other areas appears to affect the perception of recovery. In terms of the VAS assessing eating and speaking abilities, physical appearance, work performance, and sleep impairment, no statistical differences were detected between the two groups.

Treatment costs

Only one article has investigated the cost-effectiveness of the MSFE from the perspective of health economics [36]. It was proved experimentally that total treatment costs in the test group were 9.99 (0.19) ten-thousand yuan while the control group spent more for MSFE than the test group, at 14.32 (2.21) ten-thousand yuan. It seemed that there were lower costs for bone grafts and collagen membranes in the test group.

Certainty of evidence

The certainty of evidence for clinical outcomes was assessed at moderate or low grade using GRADEpro GDT. Details on the strength of the available evidence are presented in Table IV. The certainty of the evidence was downgraded for the following parameters: risk of bias, inconsistency, and imprecision.

Discussion

Regarding the biological principles of MSFE without bone grafts, some scholars have speculated that the Schneiderian membrane has osteogenic potential and contains mesenchymal stem cells. Srouji *et al.* analysed the ability of Schneiderian membrane cells to undergo osteogenic differentiation using *in vitro* assays. These results suggest that alkaline phosphatase, bone morphogenic protein-2, osteopontin, osteonectin, and osteocalcin are expressed in Schneiderian membrane cells [37]. However, an increasing

Table IV. Summary of findings and strength of evidence outcomes

Certainty assessment		No. of patients					Effect		Certainty	Importance		
No. of studies	Study design	Risk of bias	Inconsistency	Indirectness	Imprecision	Other considerations	Sinus lift without graft materials for MSFE	Sinus lift with graft materials for MSFE	Relative (95% CI)	Absolute (95% CI)		
Survival rate												
11	Randomised trials	Serious	Not serious	Not serious	Not serious	None	229/234 (97.9%)	227/239 (95.0%)	RD 0.03 (-0.01 to 0.07)	30 fewer per 1,000 (from 70 fewer to 10 more)	⊕⊕⊕○ Moderate	CRITICAL
Perforation												
13	Randomised trials	Serious	Not serious	Not serious	Not serious	None	21/194 (10.8%)	15/200 (7.5%)	RD 0.03 (-0.02 to 0.09)	30 fewer per 1,000 (from 90 fewer to 20 more)	⊕⊕⊕○ Moderate	CRITICAL
ESBG												
13	Randomised trials	Not serious	Serious	Not serious	Serious	None	187	192	-	MD 1.24 lower (1.91 lower to 0.57 lower)	⊕⊕○○ Low	IMPORTANT
ISQ												
4	Randomised trials	Serious	Not serious	Not serious	Serious	None	82	85	-	MD 2.04 lower (3.66 lower to 0.41 lower)	⊕⊕○○ Low	CRITICAL
MBL												
4	Randomised trials	Not serious	Not serious	Not serious	Very serious	None	73	82	-	MD 0.05 higher (0.14 lower to 0.23 higher)	⊕⊕○○ Low	IMPORTANT

number of researchers suggest that the space created through MSFE, which provides a blood clot a stable void, may form new bone underneath the raised membrane, which is similar to the healing process of sockets [38]. Bone marrow stem cells (MSCs) in blood clots contain osteoblast progenitors, which can lead to bone formation on scaffolds in the presence of growth factors [39]. In this meta-analysis, endo-sinus bone formation was detected in two groups, although ESBG in the control group was 1.24 mm higher than that in MSFE without any bone grafts. The results of this review were consistent with Lie's opinion [12]. An animal study showed that airflow in the maxillary sinus and tension of the Schneider membrane produced pressure in the bone increment area, accelerating the absorption of the graft material. Bone materials under the Schneider membrane provide a more stable space for bone formation [40]. Considering the inadequate number of studies included in previous meta-analyses, meta-regression analyses were conducted in this review to explore the sources of high heterogeneity. This suggests that the approaches for vertical bone augmentation and follow-up period may have an impact on the interpretation of the results. Most studies were reported based on short-term follow-ups. TSA, sensitivity analyses, and assessments for publication bias indicated that short-term results from the meta-analyses were reliable in the context of ESBG.

The meta-analyses demonstrated that a higher ISQ which was evaluated more than 6 months after the sinus augmentation procedure was found in the control group. However, the TSA demonstrated that the current results from the pooled data might be the presence of false positives. There was no difference between the two groups when Fouad's study was excluded from the sensitivity analysis. However, a potential publication bias might not have been detected because of the inadequate number of articles included in this review. Schwarz discovered that chemically modified titanium surfaces promote bone regeneration [41]. In Fouad's study, the surface and type of implants used in the MSFE were not reported, in contrast to the remaining three articles. We speculate that the implant surface may influence the necessity of bone materials for MSFE. A review article reported that herbal materials can induce cell differentiation and accelerate bone regeneration, making them potential candidates for surface nanocoating on dental titanium implants [42]. Hence, more RCTs describing implant surfaces can be conducted using the MSFE with or without bone grafts in the future.

Of the five studies that reported BD, only one showed no statistical difference in BD measured after 6 months between the two groups. The other

four studies showed a higher BD in the MSFE with bone graft group. None of the studies mentioned the exclusion criteria for postmenopausal women or patients with osteoporosis before performing the trials. Borges *et al.* measured BD in a control group filled with AB under a raised sinus membrane. Compared to other bone materials, Tosta *et al.* reported that autogenous bone harvested intraorally showed a higher absorption rate for maxillary sinus grafting in humans [43]. We speculate that bone material and oestrogen levels may affect BD after MSFE, with or without bone grafts. Titanium implants may have an adverse effect on osseointegration in patients with bone metabolism defects such as osteoporosis [44]. An animal study demonstrated that grape seed extract (GSE) improved implant osseointegration in a bone model of oestrogen deficiency [45]. Multi-level factorial randomised controlled trials should be conducted to explore the use of GSE in implant dentistry.

The meta-analysis of survival rates showed no difference between the two groups (0.03, 95% CI: $-0.01-0.07$, $p = 0.17$). Both groups had high survival rates > 95% (non-grafted, 97.93%; grafted, 95.21%). Sensitivity analysis showed that the results from the pooled data with low heterogeneity were robust. On one hand, publication bias was not detected in the funnel plot or Egger's test ($p = 0.062$). In contrast, the Z-curve of TSA reached the RIS (307) and crossed the futility areas when Maximiano's study was added to the meta-analysis, which was sufficient to illustrate that there was no statistical difference in the implant survival rates between MSFE without and with bone grafts. Based on the bone engineering tissue, the bone defect area can exhibit bone formation in the presence of seed cells, scaffold materials, and the microenvironment (MSCs as seed cells, bone materials/implant as scaffold, and blood clots as microenvironment). Although a higher ESBG level was detected in the grafted group ($p = 0.0003$), it did not seem to be the pivotal factor affecting the survival rate of implants after MSFE, with or without bone grafts. Another systematic review assessed the effects of the MSFE without bone materials [12]. Lie *et al.* also reported a high survival rate for dental implants after MSFE with and without augmentation materials (98.73% and 97.92%, respectively). However, that review only evaluated MSFE using the lateral window approach for implantation and failed to estimate the sample size for the meta-analysis. We concluded that the difference between the two groups was not significant for lateral window osteotomy or transalveolar maxillary sinus floor elevation. We speculated that RBH (RBH in conventional MSFE ≤ 5 mm, RBH in non-graft MSFE > 5 mm) was not a factor influencing survival rate. It should be not-

ed that some studies were limited to 6 months. More studies with longer follow-up durations are required to evaluate the long-term survival rates of implants.

As a common complication, perforation of the Schneiderian membrane occurs in 12–44% of cases, according to a previous study [46]. Our review tried to perform a perforation analysis of MSFE with and without bone materials. Our data suggest that MSFE with or without bone materials did not have an impact on perforation of the sinus membrane ($p = 0.25$). This conclusion was confirmed by TSA, funnel plot analysis, Egger's test, and sensitivity analysis. For instance, in a study by Borges *et al.*, sinus mucosal perforations (< 2 mm) were left without a membrane or suture. Felice seals the perforation at the Bio-Oss site with an ion barrier. Hence, if perforations are appropriately handled during the MSFE procedure, there is no difference in the survival rates for implants placed in the maxillary sinus. Our conclusion is similar to Benedicta's opinion that perforations do not have a negative impact on survival rate or long-term new bone stability [47].

MBL has a significant impact on the long-term survival rate of implants. Within the first year of functioning after implant placement, the resorption of the alveolar bone around the implants did not exceed 2 mm. The mean annual absorption of the peri-implant alveolar bone after 1 year should be less than 0.2 mm. Meeting both these criteria can be considered a successful implantation [48]. In terms of the MBL in the MSFE, the results revealed no significant difference between the two groups ($p = 0.62$). The MBL in both groups met the criteria for successful implantation. Nonetheless, the inadequate sample size of the current trials may have affected the detection of potential publication bias. As TSA was not performed, the current conclusion regarding MBL has the possibility of false negatives. MBL is inevitable whether bone materials are used for MSFE or not. MBL is associated with the smoking status, occlusal load, oral hygiene, and alveolar bone quality.

Only one study compared the VAS scores for postoperative sensation between the two groups. No difference was detected between the non-grafted and grafted sides. However, the non-graft group reported fewer days of pain and sick leave than the graft group. An autogenous bone graft was used in the graft group, which may have caused swelling and pain in the donor area. Other bone materials (xenografts, allografts, and nanoparticles) have not been reported. Whether the use of these materials can achieve similar recovery as the group without bone grafts requires further investigation. One of the factors affecting postoperative recovery is operation time. Based on available evidence, there is still no clear con-

clusion between the two groups. We considered that the skill level of the surgeons would have an important effect on the surgical time.

Only one article analysed the surgical costs of the two groups from the perspective of health economics. One thing is certain: that if both groups can achieve the same long-term success rate for implants, it can reduce the cost of bone graft materials for patients. Patients will pay an additional fee if a 3D barrier guide is used as a tent in non-graft MSFE. However, the various biomaterials used for MSFE vary in cost and performance depending on patient preference. On the other hand, price is not the only measurement of dental implant services. From the patient's perspective, the cost to patients depends on multiple factors. For instance, many patients from towns are unwilling to seek medication in county hospitals. They prefer driving to get medical treatments in municipal hospitals. Patients fail to take accommodation cost, time cost, and travel cost into consideration. The cost of dental implants in county hospitals is much cheaper than that in municipal hospitals at the same technical level. Consequently, the hierarchical medical system is of benefits for patients, doctors and the government. From the perspective of accounting, economics and management decision makers, the price of a medical service consists of direct and indirect costs. The direct cost is the sum of personnel wages and the consumption of sanitary materials. A dentist's salary is related to his market value. The market value of an experienced expert and a less experienced doctor must be different. The indirect cost refers to the consumption for production. For example, hospital administration, logistics managers, and medical and technical departments share the costs of the department. We plan to conduct further studies on various graft materials in MSFE. Cost-effectiveness analyses of biomaterials in MSFE can be performed in the future.

The limitations of this review are as follows. (1) The strength of evidence outcomes is low and moderate, and conclusions are likely to change in the future. (2) Most studies draw conclusions based on short-term follow-ups. (3) Some RCTs included in this review failed to control for confounding factors. (4) All results from meta-analyses need to be interpreted with caution. As a clinical dentist, personal experience, medical evidence, and individual differences should be combined to serve patients.

In conclusion, this systematic review revealed that both non-graft and grafted MSFE had high implant survival rates. Although the grafted MSFE had a higher ESBG, this did not seem to be a determining factor for implant survival. According to the available evidence, there was no significant

difference in the MBL and perforation rates between the two groups.

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Conflict of interest

The authors declare no conflict of interest.

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